# Sparse neural networks for forward and inverse estimation

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### Motivation

Goal: build self-learning neural network based meta-models



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Inverse problem

Given noisy observations

 $y_{n+1} = Hz_{n+1} + \varepsilon_{n+1}$ 

find the uknown  $\boldsymbol{z}$  and  $\boldsymbol{q}$  modelled by

$$z_{n+1} = G(z_n, q) + \eta_n, \quad n \in \mathbb{N}$$

in predefined time interval [0, T].

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- the operator  $H \in \mathcal{L}(\mathbb{R}^d \mapsto \mathbb{R}^m)$
- $(\varepsilon_n), (\eta_n)$  are i.i.d sequences

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We may distinguish: state and parameter estimation problems.

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For an efficient estimation of the posterior we need two steps:

• Forecast (prediction, uncertainty propagation) step

Map 
$$\phi: q_f(\omega) \mapsto y_{n,f}(\omega)$$

Assimilation (update) phase

$$\operatorname{Map}\,\varphi: y_{n,f}(\omega) \mapsto q_f(\omega)$$

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Both of these maps can be approximated by neural networks.

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## Time-dependent neural networks

In this talk we will look at:

- standard recurrent neural network
- standard long-short term memory network <sup>a</sup>



#### <sup>a</sup>pics @towardsdatascience

### Time-dependent neural network

Let be given the nonlinear dynamical system described by

$$\dot{\boldsymbol{x}}(t) = \boldsymbol{f}(\boldsymbol{x}(t), \boldsymbol{x}_t, \boldsymbol{u}(t), t)$$

in which  $oldsymbol{x} \in \mathbb{R}^d$  is the state of the system,  $oldsymbol{u} \in \mathbb{R}^d$  is the input

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For the discrete delay one has:

$$\dot{\boldsymbol{x}}(t) = \boldsymbol{f}(\boldsymbol{x}(t), \boldsymbol{x}(t-\tau_1), ..., \boldsymbol{x}(t-\tau_m), \boldsymbol{u}(t), t)$$

in which  $\tau_1 > \cdots > \tau_m \ge 0$  denotes the memory of the system.

### Time dependent neural network

In a special linear case [Sherstinsky,2020] the previous system can be decoupled to

$$\dot{\boldsymbol{x}}(t) = \boldsymbol{A}\boldsymbol{x}(t) + \boldsymbol{B}\boldsymbol{h}(t-\tau_0) + \boldsymbol{C}\boldsymbol{u}(t) + \boldsymbol{a}, \quad \boldsymbol{h} = \boldsymbol{g}(\boldsymbol{x}(t-\tau_0))$$

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in which g is a nonlinear, saturating, and invertible function of a state. Taking  $\tau_0 = \Delta t$ , and after the discretization by the implicit Euler technique, one may rewrite the previous system by

$$oldsymbol{x}_n = oldsymbol{W}_x oldsymbol{x}_{n-1} + oldsymbol{W}_h oldsymbol{h}_{n-1} + oldsymbol{W}_u oldsymbol{u}_n + oldsymbol{b}$$
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in which

$$W_x = (I - \Delta t A)^{-1}, \quad W_h = \Delta t W_x B$$
  
 $W_u = \Delta t W_x C, \quad b = \Delta t W_s a.$ 

B. Rosić (AMDA, UTwente)

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$$a_{ii} << 0, a_{i \neq j} = 0$$
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• for  $oldsymbol{B} = oldsymbol{B}^T, oldsymbol{B} = oldsymbol{V}_B oldsymbol{\Lambda}_B oldsymbol{V}_B^T$  then the stability comes from

$$\boldsymbol{W}_s := -\boldsymbol{A}^{-1}\boldsymbol{B} = -(\boldsymbol{V}_B^T\boldsymbol{A}^{-1})(\boldsymbol{V}_B\boldsymbol{\Lambda}_B).$$

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$$\boldsymbol{W}_s := -\boldsymbol{A}^{-1}\boldsymbol{B} = -(\boldsymbol{V}_B^T\boldsymbol{A}^{-1})(\boldsymbol{V}_B\boldsymbol{\Lambda}_B).$$

Thus, sufficient condition for stability is

$$0 < \lambda_i (|a_{ii}|)^{-1} < 1$$
, i.e.  $0 < \lambda_i < |a_{ii}|$ 

#### Reccurent neural cell

As  $oldsymbol{W}_x$  vanishes, the previously described dynamical system becomes

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This matches the definition of the standard recurrent neural cell (i.e. the evolution of a hidden state)

$$\boldsymbol{h}_n = \boldsymbol{g}(\boldsymbol{W}_h \boldsymbol{h}_{n-1} + \boldsymbol{W}_u \boldsymbol{u}_n + \boldsymbol{b})$$

with the output (observable) defined as

$$\boldsymbol{y}_n = Y(\boldsymbol{h}_n, \boldsymbol{w}_y)$$

in which Y is possibly a nonlinear observation operator.



#### Reccurent neural network

Collecting

$$\boldsymbol{w} := \{ \boldsymbol{W}_h, \boldsymbol{W}_u, \boldsymbol{b}, \boldsymbol{w}_y \}$$

and applying recurrency in the time interval  $[0, \Delta t, ..., n\Delta t]$ , one can further introduce a neural network (NN) as a composition of n functions

$$\boldsymbol{g}_n(\boldsymbol{h}, \boldsymbol{u}, \boldsymbol{w}) := \boldsymbol{g}(\boldsymbol{W}_h \boldsymbol{h}_{n-1} + \boldsymbol{W}_u \boldsymbol{u}_n + \boldsymbol{b})$$

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such that NN reads

$$oldsymbol{F}(oldsymbol{h},oldsymbol{u},oldsymbol{w}):=(oldsymbol{Y}_n\circoldsymbol{g}_{n-1}\circoldsymbol{g}_{n-2}\circ...\circoldsymbol{g}_1)(oldsymbol{h},oldsymbol{u},oldsymbol{w}).$$

Here,

$$Y_n(\boldsymbol{h}, \boldsymbol{u}, \boldsymbol{w}) := Y(\boldsymbol{h}_n, \boldsymbol{w}_y).$$

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#### Reccurent neural network

As all cells share weights, we have significant reduction of the parametrisation compared to the feedforward network.



The goal is to estimate  $oldsymbol{w}$  given data  $(oldsymbol{y}_i)_{i=1,n}$  such that

$$oldsymbol{w}^* = rg\minoldsymbol{J}(oldsymbol{w}), \quad oldsymbol{J}(oldsymbol{w}) \coloneqq \sum_{i=1}^n rac{1}{2} \langle oldsymbol{y}_i - \hat{oldsymbol{y}}_i(oldsymbol{w}), oldsymbol{y}_i - rac{\hat{oldsymbol{y}}_i(oldsymbol{w})}{NN} 
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$$\boldsymbol{w} = \boldsymbol{w} - lpha \frac{\partial \boldsymbol{J}}{\partial \boldsymbol{w}}$$

then requires estimation of the gradient J that depends on

$$\|\boldsymbol{W}_h\|^{\ell-n}\|\boldsymbol{g}'(\boldsymbol{z})\|^{\ell-n},$$

and thus on the properties of both  $\|\boldsymbol{W}_h\|$  and  $\|\boldsymbol{g}'(\boldsymbol{z})\|$ .

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one may have the following scenarios

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• If all  $0 < \lambda_i(\boldsymbol{W}_h) < 1$  then  $\||\boldsymbol{W}_h\| < 1$ , and if  $\|\boldsymbol{g}'(\boldsymbol{z})\| < 1$  then the gradient vanishes.

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- If all  $0 < \lambda_i(\boldsymbol{W}_h) < 1$  then  $\||\boldsymbol{W}_h\| < 1$ , and if  $\|\boldsymbol{g}'(\boldsymbol{z})\| < 1$  then the gradient vanishes.
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- If any  $\lambda_i(\boldsymbol{W}_h) > 1$  then the term  $\|\boldsymbol{W}_h\|$  will exponentially grow, and thus two scenarios:
  - If ||g'(z)|| = 0 (the flat regions of the activation function), then the gradient vanishes
  - If ||g'(z)|| ≠ 0 (quasi-linear regions of the activation function), then the gradient explodes.

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Looking at

$$\||oldsymbol{W}_h\|^{\ell-n}\|oldsymbol{g}'(oldsymbol{z})\|^{\ell-n}$$

one may have the following scenarios

- If all  $0 < \lambda_i(\boldsymbol{W}_h) < 1$  then  $\||\boldsymbol{W}_h\| < 1$ , and if  $\|\boldsymbol{g}'(\boldsymbol{z})\| < 1$  then the gradient vanishes.
- If any  $\lambda_i(\boldsymbol{W}_h) > 1$  then the term  $\|\boldsymbol{W}_h\|$  will exponentially grow, and thus two scenarios:
  - If ||g'(z)|| = 0 (the flat regions of the activation function), then the gradient vanishes
  - If ||g'(z)|| ≠ 0 (quasi-linear regions of the activation function), then the gradient explodes.

Thus, the RNNs suffer from the so-called gradient problem when used in long term integration.

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### Long-short term memory cell

To make the system robust one may generalize previous equation to

$$\boldsymbol{x}_n = \boldsymbol{g}_{cx}(n) \odot (\boldsymbol{W}_x \boldsymbol{x}_{n-1}) + \boldsymbol{g}_c(n) \odot \boldsymbol{g}(\boldsymbol{s}_n)$$

in which

$$\boldsymbol{s}_n := \boldsymbol{W}_h \boldsymbol{v}_{n-1} + \boldsymbol{g}_{cu}(n) \odot \boldsymbol{W}_u \boldsymbol{u}_n + \boldsymbol{b}$$

and

$$\boldsymbol{v}_{n-1} := \boldsymbol{g}_{ch}(n) \odot \boldsymbol{h}_{n-1}.$$

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#### Long-short term memory cell

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and

$$\boldsymbol{v}_{n-1} := \boldsymbol{g}_{ch}(n) \odot \boldsymbol{h}_{n-1}.$$

Here, controls are continuous, differentiable, monotonically increasing functions that map the domain  $(-\infty,\infty)$  into the range (0,1) (e.g. logistic function), i.e.  $0 \leq g_{cx}(n), g_{cu}(n), g_{c}(n) \leq 1$ .

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$$\boldsymbol{v}_{n-1} := \boldsymbol{g}_{ch}(n) \odot \boldsymbol{h}_{n-1}.$$

Here, controls are continuous, differentiable, monotonically increasing functions that map the domain  $(-\infty,\infty)$  into the range (0,1) (e.g. logistic function), i.e.  $0 \leq g_{cx}(n), g_{cu}(n), g_{c}(n) \leq 1$ . Taking  $W_x = I$  one obtains

$$\boldsymbol{x}_n = \boldsymbol{g}_{cx}(n) \odot \boldsymbol{x}_{n-1} + \boldsymbol{g}_{cu}(n) \odot \boldsymbol{g}(\boldsymbol{s}_n)$$

which is a core constituent of the set of formulas defining the cell of the Vanilla LSTM network [ Hochreiter and Schmidhuber,1997].

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## Long-short term memory cell

Thus,

$$oldsymbol{x}_n = oldsymbol{g}_{cx}(n) \odot oldsymbol{x}_{n-1} + oldsymbol{g}_{cu}(n) \odot oldsymbol{g}(oldsymbol{s}_n)$$
 $oldsymbol{h}_n = g(oldsymbol{x}_n)$ 

in which we choose

$$g_{cx}(n) = g_a(\hat{W}_x x_n + \hat{W}_h h_{n-1} + b_{cx})$$
$$g_{cu}(n) = g_a(\tilde{W}_x x_n + \tilde{W}_h h_{n-1} + b_{cu})$$
$$g_c(n) = g_a(\bar{W}_x x_n + \bar{W}_h h_{n-1} + b_c)$$

This matches the definition of the standard LSTM cell with the output (observable) defined as

$$\boldsymbol{y}_n = Y(\boldsymbol{h}_n, \boldsymbol{w}_y)$$

in which Y is a possibly nonlinear observation operator.

B. Rosić (AMDA, UTwente)

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## Offline gradient descent

Collecting all unknown parameters to

$$\boldsymbol{w} := \{ \hat{\boldsymbol{W}}_x, \hat{\boldsymbol{W}}_h, \tilde{\boldsymbol{W}}_x, \tilde{\boldsymbol{W}}, ... \}$$

the goal is to estimate  $oldsymbol{w}$  given data  $(oldsymbol{u}_n,oldsymbol{y}_n)$  such that

$$oldsymbol{w}^* = rg\minoldsymbol{J}(oldsymbol{w}), \quad oldsymbol{J}(oldsymbol{w}) := \sum_{i=1}^n rac{1}{2} \langle oldsymbol{y}_i - \hat{oldsymbol{y}}_i(oldsymbol{w}), oldsymbol{y}_i - \hat{oldsymbol{y}}_i(oldsymbol{w}) 
angle$$

is minimized by using the gradient based approach

$$w = w - \alpha \frac{\partial J}{\partial w}.$$

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## But, we dont have sparsity...

and cannot include noise in data, or in the input...



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## Stochastic RNN formulation

Let the unknown weights  $oldsymbol{w}$  be modelled as uncertain, i.e.

 $\boldsymbol{w}(\omega_w) \in L_2(\Omega_w, \mathfrak{F}_w, \mathbb{P}_w)$ 

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## Stochastic RNN formulation

Let the unknown weights  $\boldsymbol{w}$  be modelled as uncertain, i.e.

$$\boldsymbol{w}(\omega_w) \in L_2(\Omega_w, \mathfrak{F}_w, \mathbb{P}_w)$$

such that the RNN cell based dynamical system becomes stochastic

$$\hat{\boldsymbol{x}}_n(\omega_w) = \boldsymbol{W}_h(\omega_w)\hat{\boldsymbol{h}}_{n-1}(\omega_w) + \boldsymbol{W}_u(\omega_w)\boldsymbol{u}_n(\omega_w) + \boldsymbol{b}(\omega_w)$$
 $\hat{\boldsymbol{h}}_n(\omega_w) = \boldsymbol{g}(\hat{\boldsymbol{x}}_n(\omega_w))$ 

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such that the RNN cell based dynamical system becomes stochastic

$$\hat{x}_n(\omega_w) = W_h(\omega_w)\hat{h}_{n-1}(\omega_w) + W_u(\omega_w)u_n(\omega_w) + b(\omega_w)$$
  
 $\hat{h}_n(\omega_w) = g(\hat{x}_n(\omega_w))$ 

with the output (observable) defined as

$$\hat{\boldsymbol{y}}_n(\omega_w) = Y(\hat{\boldsymbol{h}}_n(\omega_w), \boldsymbol{w}_y(\omega_w)) + \boldsymbol{\varepsilon}_n(\omega_{\varepsilon})$$

in which  $\varepsilon_n(\omega_{\varepsilon})$  is the prediction of the cell-modelling and/or observation error, here assumed to be independent of  $w(\omega_w)$ .

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# Stochastic RNN cell: forward pass

Given  $oldsymbol{w}(\omega_w), oldsymbol{arepsilon}(\omega_arepsilon)$  and

$$\Omega := \Omega_w \times \Omega_{\varepsilon}, \mathfrak{F} := \sigma(\mathfrak{F}_w \times \mathfrak{F}_{\varepsilon}), \mathbb{P} = \mathbb{P}_w \mathbb{P}_{\varepsilon}$$

we may estimate the predicted values

$$\hat{\boldsymbol{x}}_{n}(\omega) = \boldsymbol{W}_{h}(\omega)\hat{\boldsymbol{h}}_{n-1}(\omega) + \boldsymbol{W}_{u}(\omega)\boldsymbol{u}_{n}(\omega) + \boldsymbol{b}(\omega)$$
$$\hat{\boldsymbol{h}}_{n}(\omega) = \boldsymbol{g}(\hat{\boldsymbol{x}}_{n}(\omega))$$
$$\hat{\boldsymbol{y}}_{n}(\omega) = Y(\hat{\boldsymbol{h}}_{n}(\omega), \boldsymbol{w}_{y}(\omega)) + \boldsymbol{\varepsilon}_{n}(\omega)$$

with one of the following methods

- sampling (e.g. Monte Carlo, quasi-Monte Carlo, etc.)
- approximation based methods (e.g. kernel methods, Gaussian mixture, etc.)

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# **Bayesian RNN**

Given observation

$$\boldsymbol{y}_n = Y(\boldsymbol{h}_n(\boldsymbol{w}), \boldsymbol{w}) + \boldsymbol{\varepsilon}(\hat{\omega}),$$

one may estimate the unknown weights  $m{w}$  by using Bayes rule

 $p(\boldsymbol{w}|\boldsymbol{y}_n) \propto p(\boldsymbol{y}_n|\boldsymbol{w})p(\boldsymbol{w})$ 

in which  $p(y_n|w)$  denotes the likelihood, and p(w) is the a priori distribution.

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## **Bayesian RNN**

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$$\boldsymbol{y}_n = Y(\boldsymbol{h}_n(\boldsymbol{w}), \boldsymbol{w}) + \boldsymbol{\varepsilon}(\hat{\omega}),$$

one may estimate the unknown weights w by using Bayes rule

$$p(\boldsymbol{w}|\boldsymbol{y}_n) \propto p(\boldsymbol{y}_n|\boldsymbol{w})p(\boldsymbol{w})$$

in which  $p(\pmb{y}_n|\pmb{w})$  denotes the likelihood, and  $p(\pmb{w})$  is the a priori distribution.

Assuming that all activation functions and observation are linear, and

$$p(\boldsymbol{w}) \sim \mathcal{N}(\boldsymbol{w}_f, \boldsymbol{C}_w), \quad \boldsymbol{w}_f \sim \mathcal{N}(\boldsymbol{w}_f(\omega), \boldsymbol{C}_w), \quad p(\boldsymbol{\varepsilon}_n) \sim \mathcal{N}(\boldsymbol{0}, \boldsymbol{I}_{\varepsilon})$$

the Bayes's rule reduces to the regularized RNN-cost function

$$\boldsymbol{J}_{BR} := \left( \boldsymbol{J}(\boldsymbol{w}) + \frac{1}{2} \langle \boldsymbol{w} - \boldsymbol{w}_f, \boldsymbol{w} - \boldsymbol{w}_f \rangle_{\boldsymbol{C}_{\boldsymbol{w}}} + \frac{1}{2} \langle \boldsymbol{w}_f - \bar{\boldsymbol{w}}, \boldsymbol{w}_f - \bar{\boldsymbol{w}} \rangle_{\boldsymbol{C}_{\boldsymbol{w}}} \right)$$

From

$$\boldsymbol{w}^* = rg\min \boldsymbol{J}_{BR}(\boldsymbol{w})$$

the maximum aposterori estimate reads

$$oldsymbol{w}_a(\omega) = oldsymbol{w}_f(\omega) + oldsymbol{K}(oldsymbol{y}_n - \hat{oldsymbol{y}}_n(\omega))$$

in which "a" denotes a-posteriori random variable, and

$$oldsymbol{K} = oldsymbol{C}_{oldsymbol{w}(\omega),oldsymbol{y}_n(\omega)}oldsymbol{C}_{oldsymbol{y}_n(\omega)}^{-1}$$

is known as the Kalman gain. The previous formula is also known as a classical Kalman filter estimate.

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## Nonlinearity issue

However, RNN cell is violating linearity assumption:

$$\begin{split} \boldsymbol{x}_n(\omega) &= \boldsymbol{W}_h(\omega)\boldsymbol{h}_{n-1}(\omega) + \boldsymbol{W}_u(\omega)\boldsymbol{u}_n(\omega) + \boldsymbol{b}(\omega) \\ & \boldsymbol{h}_n(\omega) = \boldsymbol{g}(\boldsymbol{x}_n(\omega)) \\ & \boldsymbol{y}_n(\omega) = \boldsymbol{Y}(\boldsymbol{h}_n(\omega), \boldsymbol{w}_y(\omega)) + \boldsymbol{\varepsilon}_n(\omega) \end{split}$$

and thus one cannot use the previously described Kalman filter.

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and thus one cannot use the previously described Kalman filter.

On the other hand, estimating the full posterior using Bayes's rule:

$$p(\boldsymbol{w}|\boldsymbol{y}_n) \propto p(\boldsymbol{y}_n|\boldsymbol{w})p(\boldsymbol{w})$$

would be computationally expensive.

#### Inverse problem

Instead of estimating  $p(w|y_n)$ , estimate the conditional expectation  $\mathbb{E}(w|y_n) = \int wp(w|y_n) dw$  directly without integration.

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#### Inverse problem

Instead of estimating  $p(w|y_n)$ , estimate the conditional expectation  $\mathbb{E}(w|y_n) = \int wp(w|y_n) dw$  directly without integration.

$$\mathbb{E}(\boldsymbol{w}|\mathfrak{B}) = P_{\mathfrak{B}}(\boldsymbol{w}) = \operatorname*{arg\ min}_{\eta\in\mathscr{Q}_{\mathfrak{B}}} \|\boldsymbol{w}-\boldsymbol{\eta}\|_{\mathscr{Q}}^{2}, \quad \mathfrak{B} := \sigma(\boldsymbol{y}_{n})$$

Optimality and orthogonality conditions:

$$\begin{split} \forall \tilde{\boldsymbol{w}} \in \mathcal{Q}_{\mathfrak{B}} : \langle\!\langle \boldsymbol{w} - \mathbb{E}(\boldsymbol{w} | \sigma(\boldsymbol{y})), \tilde{\boldsymbol{w}} \rangle\!\rangle &= \\ 0 \Rightarrow \boldsymbol{w} - \mathbb{E}(\boldsymbol{w} | \sigma(\boldsymbol{y})) \in \mathcal{Q}_{\mathfrak{B}}^{\perp} \end{split}$$

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#### Inverse problem

Instead of estimating  $p(w|y_n)$ , estimate the conditional expectation  $\mathbb{E}(w|y_n) = \int wp(w|y_n) dw$  directly without integration.

$$\mathbb{E}(oldsymbol{w}|\mathfrak{B}) = P_{\mathfrak{B}}(oldsymbol{w}) = rgmin_{\eta\in\mathscr{Q}_{\mathfrak{B}}} \|oldsymbol{w}-oldsymbol{\eta}\|_{\mathscr{Q}}^2, \quad \mathfrak{B}:=\sigma(oldsymbol{y}_n)$$

Optimality and orthogonality conditions:

$$\forall \tilde{\boldsymbol{w}} \in \mathscr{Q}_{\mathfrak{B}} : \langle\!\langle \boldsymbol{w} - \mathbb{E}(\boldsymbol{w} | \sigma(\boldsymbol{y})), \tilde{\boldsymbol{w}} \rangle\!\rangle = \\ 0 \Rightarrow \boldsymbol{w} - \mathbb{E}(\boldsymbol{w} | \sigma(\boldsymbol{y})) \in \mathscr{Q}_{\mathfrak{B}}^{\perp}$$

$$\boldsymbol{w} = P_{\mathfrak{B}}\boldsymbol{w} + (I - P_{\mathfrak{B}})\boldsymbol{w}$$



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Use Doob-Dynkin lemma

 $\mathbb{E}(\boldsymbol{w}|\mathfrak{B}) = P_{\mathfrak{B}}(\boldsymbol{w}) = \boldsymbol{\varphi}(\boldsymbol{y}_n(\boldsymbol{w}))$ 

with  $\boldsymbol{\varphi} \in L_0(\mathcal{Y}; \mathcal{Q})$  such that

$$oldsymbol{w} = P_{\mathfrak{B}}oldsymbol{w} + (I - P_{\mathfrak{B}})oldsymbol{w}$$
  
 $= oldsymbol{arphi}(oldsymbol{y}_n) + (oldsymbol{w} - oldsymbol{arphi}(oldsymbol{y}_n))$ 



Then one has

$$oldsymbol{w} = \underbrace{arphi(oldsymbol{y})}_{data} + \underbrace{(oldsymbol{w} - arphi(oldsymbol{y}))}_{prior}$$

leading to [Rosic et al, 2012]

$$oldsymbol{w}_a(\omega) = oldsymbol{w}_f(\omega) + oldsymbol{arphi}(oldsymbol{y}_n) - oldsymbol{arphi}(oldsymbol{y}_n(\omega))$$

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#### Inverse problem

Given noisy data estimate the conditional expectation  $\mathbb{E}(R(w)|y_n)$  of  $\mathcal{R}$ -valued functions of w, priorly seen as vectorial RVs R(w) - in the Hilbert space  $\mathscr{R} := L_2(\Omega, \mathfrak{F}, \mathbb{P}; \mathcal{R})$ , directly without integration.

• Conditional mean:

$$R(\boldsymbol{w}) = \boldsymbol{w}$$

• Conditional covariance:

$$R(\boldsymbol{w}) = (\boldsymbol{w} - \bar{\boldsymbol{w}}) \otimes (\boldsymbol{w} - \bar{\boldsymbol{w}}), \quad \bar{\boldsymbol{w}} = \mathbb{E}(\boldsymbol{w})$$

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# Updating more than mean

Hence,

$$\mathbb{E}(R(\boldsymbol{w})|\boldsymbol{\mathfrak{B}}) = P_{\mathfrak{B}}(R(\boldsymbol{w})) = \operatorname*{arg\,min}_{\eta \in \mathscr{R}_{\mathfrak{B}}} \|R(\boldsymbol{w}) - \eta\|_{\mathscr{R}}^{2}$$

in which closed subspace

$$\mathscr{R}_{\mathfrak{B}} = L_2(\Omega, \sigma(\boldsymbol{y}_n), \mathbb{P}; \mathcal{R}), \quad \mathfrak{B} := \sigma(\boldsymbol{y}_n).$$

Optimality condition:

$$\forall \eta \in \mathscr{R}_{\mathfrak{B}} : \langle\!\langle R(\boldsymbol{w}) - \mathbb{E}(R(\boldsymbol{w}) | \mathfrak{B}), \eta \rangle\!\rangle = 0 \Rightarrow R(\boldsymbol{w}) - \mathbb{E}(R(\boldsymbol{w}) | \mathfrak{B}) \in \mathscr{R}_{\mathfrak{B}}^{\perp}$$

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# Updating more than mean

Hence,

$$\mathbb{E}(R(\boldsymbol{w})|\boldsymbol{\mathfrak{B}}) = P_{\mathfrak{B}}(R(\boldsymbol{w})) = \operatorname*{arg\,min}_{\eta \in \mathscr{R}_{\mathfrak{B}}} \|R(\boldsymbol{w}) - \eta\|_{\mathscr{R}}^{2}$$

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Optimality condition:

$$\forall \eta \in \mathscr{R}_{\mathfrak{B}} : \langle\!\langle R(\boldsymbol{w}) - \mathbb{E}(R(\boldsymbol{w}) | \mathfrak{B}), \eta \rangle\!\rangle = 0 \Rightarrow R(\boldsymbol{w}) - \mathbb{E}(R(\boldsymbol{w}) | \mathfrak{B}) \in \mathscr{R}_{\mathfrak{B}}^{\perp}$$

Using Doob-Dynkin lemma

$$\mathbb{E}(R(\boldsymbol{w})|\mathfrak{B}) = P_{\mathfrak{B}}(R(\boldsymbol{w})) = \Phi_{R(\boldsymbol{w})}(\boldsymbol{y}_n)$$

leads to [Matthies et al., 2016]

$$R(\boldsymbol{w}_a) = R(\boldsymbol{w}_f) + \Phi_{R(\boldsymbol{w})}(\boldsymbol{y}_n) - \Phi_{R(\boldsymbol{w})}(\boldsymbol{y}_n(\boldsymbol{w}_f))$$

• Exact posterior mean

$$R(\boldsymbol{w}) := \boldsymbol{w}, \quad \mathbb{E}(\boldsymbol{w}|\boldsymbol{y}_n) = \Phi_{\boldsymbol{w}}(\boldsymbol{y}_n)$$

• Exact posterior correlation

$$R(\boldsymbol{w}) = \boldsymbol{w} \otimes \boldsymbol{w}, \quad C_c := \mathbb{E}(\boldsymbol{w} \otimes \boldsymbol{w} | \boldsymbol{y}_n) = \Phi_{\boldsymbol{w} \otimes \boldsymbol{w}}(\boldsymbol{y}_n)$$

• Exact posterior covariance

$$C_p = C_c - \Phi_{\boldsymbol{w}}(\boldsymbol{y}_n) \otimes \Phi_{\boldsymbol{w}}(\boldsymbol{y}_n)$$

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In Gauss-Markov-Kalman filter

$$\boldsymbol{w}_a = \boldsymbol{w}_f + \varphi(\boldsymbol{y}_n) - \varphi(\boldsymbol{y}_{n,f}), \quad \tilde{\boldsymbol{w}}_a = \boldsymbol{w}_f - \varphi(\boldsymbol{y}_{n,f})$$

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In Gauss-Markov-Kalman filter

$$oldsymbol{w}_a = oldsymbol{w}_f + arphi(oldsymbol{y}_n) - arphi(oldsymbol{y}_{n,f}), \quad ilde{oldsymbol{w}}_a = oldsymbol{w}_f - arphi(oldsymbol{y}_{n,f})$$

one has

$$C_{\boldsymbol{w}_a} = \mathbb{E}(\tilde{\boldsymbol{w}}_a \otimes \tilde{\boldsymbol{w}}_a | \boldsymbol{y}_n) = \mathbb{E}((\boldsymbol{w}_f - \varphi(\boldsymbol{y}_{n,f})) \otimes (\boldsymbol{w}_f - \varphi(\boldsymbol{y}_{n,f})) | \boldsymbol{y}_n)$$

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In Gauss-Markov-Kalman filter

$$oldsymbol{w}_a = oldsymbol{w}_f + arphi(oldsymbol{y}_n) - arphi(oldsymbol{y}_{n,f}), \quad ilde{oldsymbol{w}}_a = oldsymbol{w}_f - arphi(oldsymbol{y}_{n,f})$$

one has

$$C_{\boldsymbol{w}_a} = \mathbb{E}(\tilde{\boldsymbol{w}}_a \otimes \tilde{\boldsymbol{w}}_a | \boldsymbol{y}_n) = \mathbb{E}((\boldsymbol{w}_f - \varphi(\boldsymbol{y}_{n,f})) \otimes (\boldsymbol{w}_f - \varphi(\boldsymbol{y}_{n,f})) | \boldsymbol{y}_n)$$

which is not same as

$$C_p = C_c - \varphi(\boldsymbol{y}_n) \otimes \varphi(\boldsymbol{y}_n), \quad C_c := \mathbb{E}(\boldsymbol{w} \otimes \boldsymbol{w} | \boldsymbol{y}_n) = \Phi_{\boldsymbol{w} \otimes \boldsymbol{w}}(\boldsymbol{y}_n)$$

In Gauss-Markov-Kalman filter

$$oldsymbol{w}_a = oldsymbol{w}_f + arphi(oldsymbol{y}_n) - arphi(oldsymbol{y}_{n,f}), \quad ilde{oldsymbol{w}}_a = oldsymbol{w}_f - arphi(oldsymbol{y}_{n,f})$$

one has

$$C_{\boldsymbol{w}_a} = \mathbb{E}(\tilde{\boldsymbol{w}}_a \otimes \tilde{\boldsymbol{w}}_a | \boldsymbol{y}_n) = \mathbb{E}((\boldsymbol{w}_f - \varphi(\boldsymbol{y}_{n,f})) \otimes (\boldsymbol{w}_f - \varphi(\boldsymbol{y}_{n,f})) | \boldsymbol{y}_n)$$

which is not same as

$$C_p = C_c - \varphi(\boldsymbol{y}_n) \otimes \varphi(\boldsymbol{y}_n), \quad C_c := \mathbb{E}(\boldsymbol{w} \otimes \boldsymbol{w} | \boldsymbol{y}_n) = \Phi_{\boldsymbol{w} \otimes \boldsymbol{w}}(\boldsymbol{y}_n)$$

Therefore, the first equation has to be corrected to

$$\boldsymbol{w}_a = \varphi(\boldsymbol{y}_n) + C_p^{1/2} C_{\boldsymbol{w}_a}^{-1/2} \tilde{\boldsymbol{w}}_a.$$

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## Still, no sparsity only noise



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# Sparsity inducing prior

In order to introduce sparsity in weights (and thus connections), we may introduce the Laplace prior [Tipping, 2001]:

$$\boldsymbol{w} \sim e^{-\|\boldsymbol{w}\|_1} \Rightarrow p(\boldsymbol{w}|\boldsymbol{\varpi}) \sim \mathcal{N}(\boldsymbol{0}, \boldsymbol{\varpi}^{-1})$$

in which  $\varpi$  is the diagonal matrix with entries  $\varpi_{ii}$  (defining precision) corresponding to the Gamma prior  $p(\varpi_{ii})$ . By marginalizing one obtains

$$p(\boldsymbol{w}) = \int p(\boldsymbol{w}|\boldsymbol{\varpi}) p(\boldsymbol{\varpi}) d\boldsymbol{\varpi}$$



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Furthermore, in  $y_n(\omega) = Y(h_n(w(\omega)), w(\omega)) + \varepsilon(\omega)$  one assumes that

$$p(\boldsymbol{\varepsilon}) \sim \mathcal{N}(\mathbf{0}, \boldsymbol{\beta}^{-1})$$

with eta also having Gamma prior, i.e. we assume eta to be unknown.

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Furthermore, in  ${\bm y}_n(\omega)=Y({\bm h}_n({\bm w}(\omega)),{\bm w}(\omega))+{\bm \varepsilon}(\omega)$  one assumes that

 $p(\boldsymbol{\varepsilon}) \sim \mathcal{N}(\mathbf{0}, \boldsymbol{\beta}^{-1})$ 

with  $\beta$  also having Gamma prior, i.e. we assume  $\beta$  to be unknown. Thus, Bayes rule reads

$$p(\boldsymbol{w}, \boldsymbol{\varpi}, \boldsymbol{\beta} | \boldsymbol{y}_n) \propto p(\boldsymbol{y}_n | \boldsymbol{w}, \boldsymbol{\varpi}, \boldsymbol{\beta}) p(\boldsymbol{w}, \boldsymbol{\varpi}, \boldsymbol{\beta})$$

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Furthermore, in  ${\bm y}_n(\omega)=Y({\bm h}_n({\bm w}(\omega)),{\bm w}(\omega))+{\bm \varepsilon}(\omega)$  one assumes that

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$$p(\boldsymbol{w}, \boldsymbol{\varpi}, \boldsymbol{\beta} | \boldsymbol{y}_n) \propto p(\boldsymbol{y}_n | \boldsymbol{w}, \boldsymbol{\varpi}, \boldsymbol{\beta}) p(\boldsymbol{w}, \boldsymbol{\varpi}, \boldsymbol{\beta})$$

The posterior is further decoupled to [Tipping, 2001]

$$p(\boldsymbol{w}, \boldsymbol{\varpi}, \boldsymbol{\beta} | \boldsymbol{y}_n) = \underbrace{p(\boldsymbol{w} | \boldsymbol{y}_n, \boldsymbol{\varpi}, \boldsymbol{\beta})}_{\text{convolution of normals } \delta(\boldsymbol{\varpi}_{MP}, \boldsymbol{\beta}_{MP})} \underbrace{p(\boldsymbol{\varpi}, \boldsymbol{\beta} | \boldsymbol{y}_n)}_{\delta(\boldsymbol{\varpi}_{MP}, \boldsymbol{\beta}_{MP})}$$

in which is again assumed that all activation functions and observation operator are linear.

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$$\begin{array}{l} \ln p(\boldsymbol{w},\boldsymbol{\varpi},\boldsymbol{\beta}|\boldsymbol{y}_n) = \underbrace{p(\boldsymbol{w}|\boldsymbol{y}_n,\boldsymbol{\varpi},\boldsymbol{\beta})}_{\text{convolution of normals}} \underbrace{p(\boldsymbol{\varpi},\boldsymbol{\beta}|\boldsymbol{y}_n)}_{\delta(\boldsymbol{\varpi}_{MP},\boldsymbol{\beta}_{MP})} \text{ the term} \\ \underbrace{p(\boldsymbol{w}|\boldsymbol{y}_n,\boldsymbol{\varpi},\boldsymbol{\beta})}_{\text{convolution of normals}} = \mathcal{N}(\boldsymbol{w}|\boldsymbol{\mu}_w,\boldsymbol{\Sigma}_w) \end{array}$$

can be estimated using the classical Kalman filter approach.

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$$\begin{array}{l} \ln p(\boldsymbol{w},\boldsymbol{\varpi},\boldsymbol{\beta}|\boldsymbol{y}_n) = \underbrace{p(\boldsymbol{w}|\boldsymbol{y}_n,\boldsymbol{\varpi},\boldsymbol{\beta})}_{\text{convolution of normals}} \underbrace{p(\boldsymbol{\varpi},\boldsymbol{\beta}|\boldsymbol{y}_n)}_{\delta(\boldsymbol{\varpi}_{MP},\boldsymbol{\beta}_{MP})} \text{ the term} \\ \underbrace{p(\boldsymbol{w}|\boldsymbol{y}_n,\boldsymbol{\varpi},\boldsymbol{\beta})}_{\text{convolution of normals}} = \mathcal{N}(\boldsymbol{w}|\boldsymbol{\mu}_w,\boldsymbol{\Sigma}_w) \end{array}$$

can be estimated using the classical Kalman filter approach.

On the other hand, the maximum point  $\delta(m{arpi}_{MP},m{eta}_{MP})$  is obtained given

$$p(\boldsymbol{\varpi}, \boldsymbol{\beta} | \boldsymbol{y}_n) \propto p(\boldsymbol{y}_n | \boldsymbol{\varpi}, \boldsymbol{\beta}) p(\boldsymbol{\varpi}) p(\boldsymbol{\beta})$$

by maximizing marginal likelihood

$$p(\boldsymbol{y}_n|\boldsymbol{arpi}, \boldsymbol{eta}) = \int p(\boldsymbol{y}_n|\boldsymbol{w}, \boldsymbol{eta}) p(\boldsymbol{w}|\boldsymbol{arpi}) d\boldsymbol{w}.$$

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## Nonlinearity

However, RNN cell is violating linearity assumption:

$$egin{aligned} oldsymbol{x}_n(\omega) &= oldsymbol{W}_h(\omega)oldsymbol{h}_{n-1}(\omega) + oldsymbol{W}_u(\omega)oldsymbol{u}_n(\omega) + oldsymbol{b}(\omega) \ && oldsymbol{h}_n(\omega) = oldsymbol{g}(oldsymbol{x}_n(\omega)) \ && oldsymbol{y}_n(\omega) = oldsymbol{Y}(oldsymbol{h}_n(\omega),oldsymbol{w}_y(\omega)) + oldsymbol{arepsilon}_n(\omega) \end{aligned}$$

and thus in

$$p(\boldsymbol{w}, \boldsymbol{\varpi}, \boldsymbol{\beta} | \boldsymbol{y}_n) = \underbrace{p(\boldsymbol{w} | \boldsymbol{y}_n, \boldsymbol{\varpi}, \boldsymbol{\beta})}_{\neq \text{convolution of normals } \delta(\boldsymbol{\varpi}_{MP}, \boldsymbol{\beta}_{MP})} \underbrace{p(\boldsymbol{\varpi}, \boldsymbol{\beta} | \boldsymbol{y}_n)}_{\boldsymbol{\delta}(\boldsymbol{\varpi}_{MP}, \boldsymbol{\beta}_{MP})}$$

is hard to estimate both of posteriors directly.

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### Nonlinear Relevance Vector Machine

The term

 $p(\boldsymbol{w}|\boldsymbol{y}_n, \boldsymbol{\varpi}, \boldsymbol{\beta})$ 

can be estimated by use of the generalized Gauss-Markov Kalman filter:

$$oldsymbol{w}_a(\omega) = oldsymbol{w}_f(\omega) + oldsymbol{arphi}(oldsymbol{y}_n) - oldsymbol{arphi}(oldsymbol{y}_n(\omega))$$

in which  $w_f(\omega) \sim \mathcal{N}(0, \varpi^{-1})$ , and similarly its covariance [Rosic, 2022, in preparation]:

$$\boldsymbol{w}_a = \varphi(\boldsymbol{y}_n) + C_p^{1/2} C_{\boldsymbol{w}_a}^{-1/2} \tilde{\boldsymbol{w}}_a.$$

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## Nonlinear Relevance Vector Machine

The term

 $p(\boldsymbol{w}|\boldsymbol{y}_n, \boldsymbol{\varpi}, \boldsymbol{\beta})$ 

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$$oldsymbol{w}_a(\omega) = oldsymbol{w}_f(\omega) + oldsymbol{arphi}(oldsymbol{y}_n) - oldsymbol{arphi}(oldsymbol{y}_n(\omega))$$

in which  $w_f(\omega) \sim \mathcal{N}(0, \varpi^{-1})$ , and similarly its covariance [Rosic, 2022, in preparation]:

$$\boldsymbol{w}_a = \varphi(\boldsymbol{y}_n) + C_p^{1/2} C_{\boldsymbol{w}_a}^{-1/2} \tilde{\boldsymbol{w}}_a.$$

On the other hand, the term

$$\underbrace{p(\boldsymbol{\varpi},\boldsymbol{\beta}|\boldsymbol{y}_n)}_{\delta(\boldsymbol{\varpi}_{MP},\boldsymbol{\beta}_{MP})}$$

is hard to estimate directly unless approximating the likelihood.

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# Approximation

In

$$\begin{split} \boldsymbol{x}_n(\omega) &= \boldsymbol{W}_h(\omega)\boldsymbol{h}_{n-1}(\omega) + \boldsymbol{W}_u(\omega)\boldsymbol{u}_n(\omega) + \boldsymbol{b}(\omega) \\ & \boldsymbol{h}_n(\omega) = \boldsymbol{g}(\boldsymbol{x}_n(\omega)) \\ & \boldsymbol{y}_n(\omega) = \boldsymbol{Y}(\boldsymbol{h}_n(\omega), \boldsymbol{w}_y(\omega)) + \boldsymbol{\varepsilon}_n(\omega) \end{split}$$

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## Approximation

In

$$\begin{split} \boldsymbol{x}_n(\omega) &= \boldsymbol{W}_h(\omega)\boldsymbol{h}_{n-1}(\omega) + \boldsymbol{W}_u(\omega)\boldsymbol{u}_n(\omega) + \boldsymbol{b}(\omega) \\ & \boldsymbol{h}_n(\omega) = \boldsymbol{g}(\boldsymbol{x}_n(\omega)) \\ & \boldsymbol{y}_n(\omega) = \boldsymbol{Y}(\boldsymbol{h}_n(\omega), \boldsymbol{w}_y(\omega)) + \boldsymbol{\varepsilon}_n(\omega) \end{split}$$

one can linearize the last two equations such that

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$$\begin{split} \boldsymbol{h}_n^{(\ell)}(\omega) &= \boldsymbol{g}^{(\ell)}(\boldsymbol{x}_n(\omega)) = \boldsymbol{J}_x \boldsymbol{x}_n(\omega) + \boldsymbol{z}_h \\ \boldsymbol{y}_n^{(\ell)}(\omega) &= \boldsymbol{Y}^{(\ell)}(\boldsymbol{h}_n(\omega), \boldsymbol{w}_y(\omega)) + \varepsilon_n(\omega) = \boldsymbol{J}_h \boldsymbol{h}_n^{(\ell)}(\omega) + \boldsymbol{z}_y + \boldsymbol{\varepsilon}_n(\omega) \end{split}$$

holds. The linearisation can be also achieved by prevously described relevance vector machine [Rosic, 2022, in preparation].

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## Gaussian approximation of the marginal likelihood

The point  $\delta(\boldsymbol{\varpi}_{MP},\boldsymbol{\beta}_{MP})$  is obtained given

$$p(\boldsymbol{\varpi}, \boldsymbol{\beta} | \boldsymbol{y}_n) \propto p(\boldsymbol{y}_n | \boldsymbol{\varpi}, \boldsymbol{\beta}) p(\boldsymbol{\varpi}) p(\boldsymbol{\beta})$$

by maximizing  $p(\boldsymbol{y}_n | \boldsymbol{\varpi}, \boldsymbol{\beta}) = \int p(\boldsymbol{y}_n | \boldsymbol{w}, \boldsymbol{\beta}) p(\boldsymbol{w} | \boldsymbol{\varpi}) d\boldsymbol{w}$  in an iterative fashion [Rosic, 2022, in preparation].

## Gaussian approximation of the marginal likelihood

The point  $\delta(\boldsymbol{\varpi}_{MP},\boldsymbol{\beta}_{MP})$  is obtained given

$$p(\boldsymbol{\varpi}, \boldsymbol{\beta} | \boldsymbol{y}_n) \propto p(\boldsymbol{y}_n | \boldsymbol{\varpi}, \boldsymbol{\beta}) p(\boldsymbol{\varpi}) p(\boldsymbol{\beta})$$

by maximizing  $p(\boldsymbol{y}_n | \boldsymbol{\varpi}, \boldsymbol{\beta}) = \int p(\boldsymbol{y}_n | \boldsymbol{w}, \boldsymbol{\beta}) p(\boldsymbol{w} | \boldsymbol{\varpi}) d\boldsymbol{w}$  in an iterative fashion [Rosic, 2022, in preparation]. After linearisation

$$\begin{split} \mathbb{E}(\boldsymbol{y}_n | \boldsymbol{w}, \boldsymbol{\beta}) &= \boldsymbol{\varPhi}^T \boldsymbol{w}, \quad \boldsymbol{C}(\boldsymbol{y}_n | \boldsymbol{w}, \boldsymbol{\beta}) = \boldsymbol{C}_{\boldsymbol{w}} \\ p(\boldsymbol{y}_n | \boldsymbol{w}, \boldsymbol{\beta}) &\approx \mathcal{N}(\mu_{\boldsymbol{w}}, \boldsymbol{C}_{\boldsymbol{w}}), \end{split}$$

the mean vector and the covariance matrix are both the functions of the weights  $\boldsymbol{w}$ .

# Gaussian approximation of the marginal likelihood

The point  $\delta(\boldsymbol{\varpi}_{MP},\boldsymbol{\beta}_{MP})$  is obtained given

$$p(\boldsymbol{\varpi}, \boldsymbol{\beta} | \boldsymbol{y}_n) \propto p(\boldsymbol{y}_n | \boldsymbol{\varpi}, \boldsymbol{\beta}) p(\boldsymbol{\varpi}) p(\boldsymbol{\beta})$$

by maximizing  $p(\boldsymbol{y}_n | \boldsymbol{\varpi}, \boldsymbol{\beta}) = \int p(\boldsymbol{y}_n | \boldsymbol{w}, \boldsymbol{\beta}) p(\boldsymbol{w} | \boldsymbol{\varpi}) d\boldsymbol{w}$  in an iterative fashion [Rosic, 2022, in preparation]. After linearisation

$$\begin{split} \mathbb{E}(\boldsymbol{y}_n | \boldsymbol{w}, \boldsymbol{\beta}) &= \boldsymbol{\varPhi}^T \boldsymbol{w}, \quad \boldsymbol{C}(\boldsymbol{y}_n | \boldsymbol{w}, \boldsymbol{\beta}) = \boldsymbol{C}_{\boldsymbol{w}} \\ p(\boldsymbol{y}_n | \boldsymbol{w}, \boldsymbol{\beta}) &\approx \mathcal{N}(\mu_{\boldsymbol{w}}, \boldsymbol{C}_{\boldsymbol{w}}), \end{split}$$

the mean vector and the covariance matrix are both the functions of the weights w. Thus, one can use the law of the total expectation to get

$$\boldsymbol{\mu} := \mathbb{E}_{p(\boldsymbol{w}|\boldsymbol{\varpi})}(\mathbb{E}(\boldsymbol{y}_n|\boldsymbol{w},\boldsymbol{\beta})) = \boldsymbol{0}$$

$$oldsymbol{C} := \mathbb{E}_{p(oldsymbol{w}|oldsymbol{arphi})}(oldsymbol{C}(oldsymbol{y}_n|oldsymbol{w},oldsymbol{eta})) + oldsymbol{C}_{p(oldsymbol{w}|oldsymbol{arphi})}(\mathbb{E}(oldsymbol{y}_n|oldsymbol{w},oldsymbol{eta}))$$

# Sparse LSTM

The complete process can be repeated for LSTM model as well [van Weg, Greve, Rosic, 2021]:

$$oldsymbol{x}_n(\omega) = oldsymbol{g}_{cx}(n,\omega) \odot oldsymbol{x}_{n-1}(\omega) + oldsymbol{g}_{cu}(n,\omega) \odot oldsymbol{g}(oldsymbol{s}_n(\omega))$$
 $oldsymbol{h}_n(\omega) = oldsymbol{g}(oldsymbol{x}_n(\omega))$ 

in which we choose

$$\begin{split} \boldsymbol{g}_{cx}(n,\omega) &= \boldsymbol{g}_{a}(\hat{\boldsymbol{W}}_{x}(\omega)\boldsymbol{x}_{n}(\omega) + \hat{\boldsymbol{W}}_{h}(\omega)\boldsymbol{h}_{n-1}(\omega) + \boldsymbol{b}_{cx}(\omega)) \\ \boldsymbol{g}_{cu}(n,\omega) &= \boldsymbol{g}_{a}(\tilde{\boldsymbol{W}}_{x}(\omega)\boldsymbol{x}_{n}(\omega) + \tilde{\boldsymbol{W}}_{h}(\omega)\boldsymbol{h}_{n-1}(\omega) + \boldsymbol{b}_{cu}(\omega)) \\ \boldsymbol{g}_{c}(n,\omega) &= \boldsymbol{g}_{a}(\bar{\boldsymbol{W}}_{x}(\omega)\boldsymbol{x}_{n}(\omega) + \bar{\boldsymbol{W}}_{h}(\omega)\boldsymbol{h}_{n-1}(\omega) + \boldsymbol{b}_{c}(\omega)) \\ \boldsymbol{y}_{n}(\omega) &= Y(\boldsymbol{h}_{n}(\omega),\boldsymbol{w}_{y}(\omega)) \end{split}$$

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Inverse problem

Given noisy data  $z \in \mathcal{Z}$ , i.e.

$$oldsymbol{z} = Z(oldsymbol{q}) + oldsymbol{\epsilon}$$

estimate the unknown  $q \in Q$ .

- $\mathcal{X} := \{\mathcal{Q}, \mathcal{Z}\}$  are Hilbert spaces with the inner product  $\langle \cdot, \cdot \rangle_{\mathcal{X}}$
- $oldsymbol{q} \in \mathcal{Q}$  is the parameter
- $Z: \mathcal{Q} \mapsto \mathcal{Z}$  is possibly nonlinear observation operator
- z are data
- $\epsilon$  are noise realisations

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By using Gauss-Markov-Kalman filter

$$\boldsymbol{q}_a(\omega) = \boldsymbol{q}_f(\omega) + \boldsymbol{\varphi}(\boldsymbol{z}_m) - \boldsymbol{\varphi}(\boldsymbol{y}_f(\omega))$$

we may distinguish two steps [van Dijk, Hakvoort, Rosic, 2022]:

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• Forecast (prediction, uncertainty quantification) step

Map 
$$\boldsymbol{\phi}: \boldsymbol{q}_f(\omega) \mapsto \boldsymbol{y}_f(\omega)$$

• Assimilation (update) phase

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Both of these maps can be approximated by sparse NNs such that:

$$oldsymbol{q}_a(\omega) = oldsymbol{q}_f(\omega) + oldsymbol{arphi}_{sNN}(oldsymbol{z}_m) - oldsymbol{arphi}_{sNN}(oldsymbol{\phi}_{sNN}(oldsymbol{q}_f(\omega)) + oldsymbol{\epsilon}(\omega)$$

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### Numerical example



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# Convergence



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Sparsity



Sparsity of a) Dense layer, b) LSTM cell

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Sparse NNs

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#### Comparison to point estimate

$$\mathbf{R}^{2} = 1 - \frac{\sum_{i}^{m} \sum_{j}^{n_{b}} \left( \boldsymbol{y}_{ij} - \frac{1}{n_{b}} \sum_{n_{b}} (\boldsymbol{y}_{i})^{*} \right)^{2}}{\sum_{i}^{m} \sum_{j}^{n_{b}} \left( \boldsymbol{y}_{ij} - \frac{1}{n_{b}} \sum_{n_{b}} \boldsymbol{y}_{i} \right)^{2}},$$

	$n_m$	epochs $[-]$	time $[s]$	time per epoch $[s]$	$R^{2}[-]$
Point estimate LSTM	16	4000	189	0.047	0.994
	32	4000	189	0.048	0.993
	64	4000	190	0.048	0.994
	128	4000	191	0.048	0.996
ARD-LSTM	16	779	153	0.20	0.993
	32	541	151	0.28	0.995
	64	497	317	0.64	0.998
	128	434	1403	3.23	0.998

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# Expected improvement



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# Identification

$$\begin{split} \frac{\mathrm{d}x}{\mathrm{d}t} &= -\sigma x + \sigma y, \\ \frac{\mathrm{d}y}{\mathrm{d}t} &= \rho x - xz - y, \\ \frac{\mathrm{d}z}{\mathrm{d}t} &= xy - \beta z. \end{split}$$

$$\begin{split} \mathbf{x}_0 &= [1.508870, -1.531271, 25.46091] \\ \mathbf{q} &= [\sigma, \rho, \beta] = [10, 28, 8/3] \\ \mathbf{q}(\omega) &\sim \mathcal{U}(\mathbf{q}_{min}, \mathbf{q}_{max}), \\ \mathbf{\mu}_{\mathbf{x}_0} &= \mathbf{x}_0, \qquad \mathbf{\sigma}_{\mathbf{x}_0}^2 = [2, 2, 2], \\ \mathbf{x}_0(\omega) &\sim \mathcal{N}(\mathbf{\mu}_{\mathbf{x}_0}, \mathbf{\sigma}_{\mathbf{x}_0}^2 I), \\ \mathbf{q}_{min} &= [1, 1, 1], \qquad \mathbf{q}_{max} = [30, 44.8, 5.3]. \end{split}$$



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## State Identification



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## State Identification



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## Parameter Identification



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# Conclusion

Currently done:

- Neural networks (NN) can be represented as delayed differential equations
- Classical training is reqiring more data due to higher parametrisation
- Sparse training using relevance vector machine is only for linear case
- We suggest nonlinear releveance vector machine and apply on NN

To be done:

- study the requirements for convergence and stability
- extend this with the model reduction techniques

### Thank you: any questions?



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